

Survey Paper

Take a bunch of local random decisions that converge to a global state that adheres to a required distribution.

#P $\rightarrow$ class of hard counting problems.

MC sampling: Allows us to approximate answers to these sort of problems.

Consider a simple MC for Indep. Set.

i) choose $v \in V$

ii) if $v \in I_t$ then move to state $I_t - \{v\}$.

iii) Else if $v \notin I_t$ and $\{v\} \cup I_t$ is a valid indep. set then move to $I_t \cup \{v\}$

iv) Else don't move.

Looking at the transition prob-

$$P(I, I') = \begin{cases} 1/n, & \text{if } d(I, I') = 1 \\ 0, & \text{if } d(I, I') > 1 \\ 1 - \sum_{J \neq I} P(I, J), & I = I' \end{cases} \quad \begin{matrix} \nearrow \text{Hamming Dist} \end{matrix}$$

Definition: MC is ergodic:

i) irreducible: $\forall x, y \exists t \text{ s.t. } p^t(x, y) > 0$.

(at some point I can reach anywhere from anywhere with non-zero probability.)

ii) aperiodic: $\forall x, y \gcd(\{t \mid p^t(x, y) > 0\}) = 1$.

(there is no bipartition, a bipartition would cause one side of the cut to be non zero at odd time steps and

the other side at even time steps.)

Notice that the indep set chain is both. Particularly, the self-loop makes sure that bipartition is impossible.

This can work on any chain - add self loops with some const prob and it will be aperiodic (however there may be a small, constant cost to the runtime).

lazy chain: $\frac{1}{2}$ prob of self loop at each state.

Main Lemma: Any finite, ergodic MC converges to a unique stationary distribution π , i.e. $\forall x, y$ as $t \rightarrow \infty$
 $p^t(x, y) = \pi(y)$.

Detailed Balance:

For an ergodic MC with trans prob P , if $\pi': \Omega \rightarrow [0, 1]$ is any function s.t.

$$\pi'(x) \cdot P(x, y) = \pi'(y) \cdot P(y, x) \text{ s.t.}$$

$\sum_{x \in \Omega} \pi'(x) = 1$ then π' is the unique stationary

distribution.

chains satisfying detailed balance: time reversible.

Notice that for the indep set chain, $p(I, I') = p(I', I) = \frac{1}{n}$ for $I \neq I'$. $\Rightarrow$ detailed balance is satisfied as -

$$\frac{1}{n} \cdot \frac{1}{|\Omega|} = \frac{1}{n} \cdot \frac{1}{|\Omega|} \Rightarrow \pi'(x) = \frac{1}{|\Omega|}$$

uniform distribution

Natural Questions -

- i) Can I choose a more complicated distribution.
- ii) How long does it take to converge.

Metropolis's Algorithm: Answer to question 1.
Starting at x repeat:

- i) Pick a neighbour y of x with prob $\frac{1}{2\Delta}$.
- ii) Move to y with prob $\min\left(1, \frac{\pi(y)}{\pi(x)}\right)$.
- iii) With remaining prob stay at x .

claim: Detail Balance is satisfied.

Verify:
$$\text{LHS} = \frac{1}{2\Delta} \min\left(1, \frac{\pi(y)}{\pi(x)}\right) \cdot \pi(x)$$

$$\text{RHS} = \frac{1}{2\Delta} \min\left(1, \frac{\pi(x)}{\pi(y)}\right) \cdot \pi(y)$$

Now if $\pi(x) > \pi(y)$

then $\text{LHS} = \frac{1}{2\Delta} \pi(y)$, $\text{RHS} = \frac{1}{2\Delta} \cdot 1 \cdot \pi(y)$
Hence equal

Similarly if $\pi(x) < \pi(y)$ also.

Going back to the IS example, suppose we want to sample with prob $\frac{\lambda^{|\mathcal{I}'|}}{Z}$ where $Z = \sum_{\mathcal{I}} \lambda^{|\mathcal{I}'|}$ is the normalizing constant.

consider $\mathcal{I}$ and $\mathcal{I}' = \mathcal{I} \cup \{v\}$ in the MC.

Here $|\mathcal{I}'| = |\mathcal{I}| + 1$

$$\Rightarrow \pi(\mathcal{I}') = \lambda \pi(\mathcal{I})$$

$$\Rightarrow p(\mathcal{I}, \mathcal{I}') = \frac{1}{2n} \min\left(\frac{\lambda \pi(\mathcal{I})}{\pi(\mathcal{I})}, 1\right) = \frac{1}{2n} \min(1, \lambda)$$

But $P(I', I) = \frac{1}{2n} \min(1, \frac{1}{\lambda})$.

Notice that Z is gone. (we don't know how to count it any way, that's a #P problem).

Again, verifying detailed balance

$$\text{LHS} = \Pi(I) \cdot \frac{1}{2n} \min(1, \lambda)$$

$$\text{RHS} = \lambda \Pi(I) \cdot \frac{1}{2n} \min(1, \frac{1}{\lambda})$$

if $\lambda < 1$, $\text{LHS} = \frac{\lambda \Pi(I)}{2n}$, $\text{RHS} = \frac{\lambda \Pi(I)}{2n} \cdot 1$.

$\lambda \geq 1$, $\text{LHS} = \frac{\Pi(I)}{2n}$, $\text{RHS} = \frac{\lambda \Pi(I)}{2n \lambda} = \frac{\Pi(I)}{2n}$

In both cases $\text{LHS} = \text{RHS}$.

Mixing Time

Time taken for an MC to converge to its stationary distribution

Defn: Total Variational Distance at time t

$$\|P^t, \Pi\|_{\text{TV}} = \max_{x \in \Omega} \frac{1}{2} \sum_{y \in \Omega} |P^t(x, y) - \Pi(y)|$$

Choose the worst source and find L1 norm. and divide by 2.

Defn: For $\epsilon > 0$, mixing time $\tau(\epsilon)$ is

$$\tau(\epsilon) = \min \{t \mid \|P^t(x, y), \Pi\|_{\text{TV}} \leq \epsilon, \forall t' \geq t\}$$

Rapidly Mixing: If $\tau(\epsilon)$ is bounded by $\text{poly}(n, \log(\epsilon))$ where n is the size of each configuration.

(for ex - #nodes in the graph for the IS chain).

For $\lambda_0, \lambda_1, \dots, \lambda_{|\Omega|-1}$ eigen values of the transition matrix s.t. $1 = \lambda_0 > |\lambda_1| \geq |\lambda_i| \forall i$ then.

Defn: Spectral Gap $\lambda_0 - |\lambda_1| = 1 - \lambda_1$.

Theorem: Let $\pi_* = \min_{x \in \Omega} \pi(x)$. For all $\epsilon > 0$,

$$a) \tau(\epsilon) \leq \frac{1}{1 - |\lambda_1|} \log\left(\frac{1}{\pi_* \epsilon}\right)$$

$$b) \tau(\epsilon) \geq \frac{|\lambda_1|}{2(1 - |\lambda_1|)} \log\left(\frac{1}{2\epsilon}\right)$$

Unfortunately this neat trick isn't useful when the transition prob matrix is unavailable. i.e. when the state space is exponentially large.

Coupling

Defn: A MC on $\Omega \times \Omega$ defining $(X_t, Y_t)_{t=0}^{\infty}$ s.t.

- i) X_t and Y_t are copies of M
- ii) If $X_t = Y_t$ then $X_{t+1} = Y_{t+1}$.

i.e. i) Each when viewed individually simulates the original chain.
 ii) Once they both agree, they agree from that time forward.

Defn: For initial states x, y

$$T^{x,y} = \min \{t \mid X_t = Y_t \text{ for } X_0 = x, Y_0 = y\}$$

i.e. the first time they agree.

Defn: Coupling Time $T = \max_{x,y} E[T_{x,y}]$

Thm: $T(e) \leq \lceil T e \ln e^{-1} \rceil$

Toy Example: Sampling a point from a boolean hypercube.
The natural flip a bit MCMC won't work since it's bipartite (odd and even # bits set). So a modified chain -

MCcube

starting at $x_0 = (0, 0, \dots, 0)$ do

i) Pick $(i, b) \in \{1, \dots, n\} \times \{0, 1\}$

ii) $x^{t+1} = x^t$ with $x^t[i] = b$.

This gives us transition prob $P(x, y) = \begin{cases} \frac{1}{2n}, & d(x, y) = 1 \\ \frac{1}{2}, & x = y \\ 0, & \text{o/w} \end{cases}$

Now if you want to couple, start from two vertices x_0 and y_0 update these with the same (i, b) .

$\Rightarrow$ If we have updated each bit at least once, by pigeonhole principle they must have collided.

Apparently this converges in $O(n \log n)$. due to coupon collector

$\Rightarrow$ Coupling Time: $n \log n$

$$\Rightarrow \text{Mixing Time } T(\epsilon) = n \frac{\ln n}{\ln 2} e \ln \epsilon^{-1} \\ = O(n \ln(n \epsilon^{-1}))$$

Coupling proofs may get more complicated if the distance b/w configurations increases. For the hypercube this distance is non decreasing (because the update is the same). When it can increase, the argument doesn't work directly, but if in expectation it decreases, we can use some theory of random walks to show an $n \log n$ coupling time.

Path Coupling

Finding expected change in distance is not straightforward.

Bubley and Dyer showed that it is sufficient to consider only pairs of configurations that are close.

i.e. consider a small set $U \subseteq \Omega \times \Omega$ that are close in some distance metric, ϕ . For ex- pairs of configs with hamming distance 1. Now suppose expected change in distance is decreasing for all pairs in U .

To reason about some arbitrary $X, Y \in \Omega$. Let $z_0 = X \dots z_r = Y$ define a shortest path from X to Y of length r , where $(z_i, z_{i+1}) \in U$.

$\Rightarrow \phi(X, Y) = \sum_{i=0}^{r-1} \phi(z_i, z_{i+1})$. Each of these individual decrease in expectation. Then by linearity $\phi(X, Y)$ also decreases in expectation.

Formally: let ϕ be an integer valued metric defined on $\Omega \times \Omega$ which takes values in $\{0, \dots, B\}$. Let U be a subset of $\Omega \times \Omega$ such that for s.t. for all (x_t, y_t) in $\Omega \times \Omega$ there exists a path $z = z_0, \dots, z_n$ s.t. $(z_i, z_{i+1}) \in U$. and

$$\sum_{i=0}^{n-1} \phi(z_i, z_{i+1}) = \phi(x_t, y_t)$$

let M be an MC on Ω with transition matrix P .

consider a random function $f: \Omega \rightarrow \Omega$, such that $\Pr[f(x) = y] = P(x, y)$.

define a coupling of the MC by $(x_t, y_t) \rightarrow (x_{t+1}, y_{t+1}) = (f(x_t), f(y_t))$

i. if there exists $\beta < 1$ s.t. $E[\phi(x_{t+1}, y_{t+1})] \leq \beta \phi(x_t, y_t)$ for all $(x_t, y_t) \in U$ then the mixing time satisfies-

$$\tau(\epsilon) \leq \frac{\ln(B\epsilon^{-1})}{1-\beta} \quad [B \text{ is max metric}]$$

ii. if $\beta = 1$: if $\exists \alpha > 0$ s.t. $\Pr[\phi(x_{t+1}, y_{t+1}) \neq \phi(x_t, y_t)] \geq \alpha$. For all x_t, y_t then

$$\tau(\epsilon) \leq \left\lceil \frac{eB^2}{\alpha} \right\rceil \lceil \ln \epsilon^{-1} \rceil$$

Glauber Dynamics

1. with prob = $1/2$ do nothing.
2. $(v, c) \in_R V \times \{1, \dots, K\}$

3. If v recolored with c is valid recolor else do nothing.

Transition matrix is symmetric $\Rightarrow$ converges to uniform distribution over all valid k -colorings.

Let's look at coupling here. Let x_0, y_0 be two arbit, valid k -colorings. Let's update both with the same update (v, c) .

Thm: $T(E) = O(n \ln(n e^{-1}))$ for an n -vertex graph with $k \geq 3\Delta + 1$.

Proof: Let x_0, y_0 be starting configurations. At each t , choose $(v, c) \in R$ and update both x_t and y_t as defined.

Let $\phi: \Omega \times \Omega \rightarrow \mathbb{Z}$ be the Hamming metric. i.e. the min no. of colors to flip to reach one config from another. Here notice that the max possible shortest distance $B = 2n$ (not sure - won't $n+1$ be enough?).

Let U be the set of pairs of colorings at $\phi = 1$. Now we are interested in change in $\phi(r, s), (r, s) \in U$.

Suppose w is the vertex that differs in r and s ,

i) If $w = v$, neighborhood is same, hence either distance decreases (if both accept the c) or remains same. There are at least $k - \Delta$ colors that will be accepted.

$$\Rightarrow \mathbb{E}[\Delta\phi] \leq -1 \times \frac{(k-\Delta)}{kn} + 0 \times \left(1 - \frac{(k-\Delta)}{kn}\right)$$

$$\leq -\frac{(k-\Delta)}{kn}.$$

ii) $(u, v) \in E$: If they both accept or reject the distance is unchanged. However if $c = \chi(v)$ for either, the distance will increase by 1.

$$\Rightarrow \mathbb{E}(\Delta\phi) \leq \frac{2}{kn} \text{ for fixed } \omega.$$

$$\Rightarrow \text{for all possible } \omega, \mathbb{E}(\Delta\phi) \leq \frac{\Delta^2}{kn}$$

iii) $u \neq v$ and $(u, v) \notin E$: Here the change in dist is always 0 since both will either accept or reject.

$$\Rightarrow \mathbb{E}[\Delta\phi] \leq \frac{1}{kn} (-k + \Delta + 2\Delta) = \frac{3\Delta - k}{kn}$$

$$\Rightarrow \text{if } k > 3\Delta, \mathbb{E}(\Delta\phi) < 1$$

$$\Rightarrow T(\epsilon) \leq \frac{\ln 2n\epsilon^{-1}}{1 - \frac{3\Delta - k}{kn}}$$

For tight bound let $k = 3\Delta + 1$

$$\Rightarrow T(\epsilon) \leq \frac{\ln n\epsilon^{-1}}{1 - \frac{1}{kn}}$$

$$= O(\ln \ln(n\epsilon^{-1}))$$

Extensions

→ choosing better coupling.

Glauber Dynamics can actually be proven to converge with the above line even if $k > 2\Delta$. This is achieved by looking carefully at case ii.

→ changing the Markov chain.

possibly add extra moves or update more than one node

at a time. Anything that makes analyzing these distances more efficient. In fact, it is proved that rapid mixing of some modified chain implies rapid mixing of the original Glauber dynamics.

→ Macromoves: Add a "warm up" or "burn in" period that randomizes the configurations before each step.

Vigoda and Hayes have extended this concept to work during the coupling phase itself.

Flows

intuition: slow mixing time is caused by some sort of bottleneck $\Rightarrow$ exponentially small cut.

Min Cut

conductance of a chain -

For $S \subset \Omega$, let

$$\phi_S = \sum_{\substack{x \in S, \\ y \notin S}} \frac{Q(x, y)}{\pi(S)}$$

where $Q(x, y) = \pi(x) \cdot p(x, y)$. This can be thought of as the capacity of edge (x, y) .

And $\pi(S) = \sum_{x \in S} \pi(x)$ is the weight of the cut set. (weight of edges crossing the cut).
(I don't see how).

$$\text{Conductance } \Phi = \min_{S | \pi(S) \leq \frac{1}{2}} \phi_S$$

Intuitively - low conductance $\Rightarrow$ there is a bad cut.

since ϕ_s can be thought of as a cut divided by $\pi(S)$. This is sort of a min cut.

Thm: For MC with conductance Φ and eig val gap $\text{Gap}(P) = 1 - |\lambda_1|$,

$$\frac{\Phi^2}{2} \leq \text{Gap}(P) \leq 2\Phi.$$

We know $\text{Gap}(P)$ relates to $\tau(\epsilon)$ as-

$$\tau(\epsilon) \leq \frac{1}{\text{Gap}(P)} \log \left(\frac{1}{\min_x \pi(x)} \right)$$

$\Rightarrow$ Small $\Phi \Rightarrow$ Small $\text{Gap} \Rightarrow$ Larger $\tau(\epsilon)$.

Max Flow

Consider a $(\Omega, E = \text{non zero transitions})$. Now min cut corresponds to a max flow

$\rightarrow$ for each pair of x, y define canonical path $\gamma_{x,y}$
 $\Gamma = \{ \gamma_{x,y} \mid x, y \in \Omega, x \neq y \}$. (some arbit path)

$\Rightarrow$ congestion $f(\Gamma) = \max_e \frac{1}{\Phi(e)} \sum_{\gamma_{x,y} \ni e} \pi(x) \pi(y)$.

i.e. an edge appearing in more paths is more congested.

$\downarrow$
some edge in the path.

Think of each path from x to y as carrying $\pi(x)\pi(y)$ units of flow. Congestion of network the max ratio of load routed through edge e divided by its capacity $Q(e)$.

Let $\bar{f} = \min_{\Gamma} f(\Gamma) \cdot l(\Gamma)$ where l is the max length of any path in Γ .

Thm: For an ergodic, reversible MC with stationary π and self loop prob $P(y,y) \geq \frac{1}{2} \quad \forall y \in \Omega$

$$T_x(e) \leq \bar{f} (\log \pi(x)^{-1} + \log e^{-1})$$

consider the hypercube chain. Choose Γ that makes paths by moving left to right and flipping bits as they mismatch.

$f(\Gamma)$: Consider an edge e .

Now $\sum \pi(x)\pi(y)$. Now $\pi(x) = \pi(y) \Rightarrow$ just have to count $\sum_{e \in \Gamma_{xy}}$ the #path containing (u,v) .

if u, v differ in the i th bit, first $i-1$ bits of u agree with the end of the path, y .

similarly last $n-i$ bits must agree with x , the start of the path.

$$\Rightarrow x = (x_1 \dots x_{i-1}, u_i, v_i \dots v_n)$$

$$y = (v_1 \dots v_i, y_{i+1}, \dots y_n)$$

$\rightarrow 2^{n-1}$ ways to assign $x_1, \dots, x_{i-1}, y_{i+1}, \dots, y_n$

$$\Rightarrow \sum_{e \in \Pi_{xy}} \pi(x) \pi(y) = \frac{1}{2^n} \cdot \frac{1}{2^n} \cdot 2^{n-1} = 2^{-(n+1)}$$

$$\text{Also, } q(e) = \frac{\overset{\pi(x)}{\uparrow}}{2^n} \cdot \frac{\overset{p(x,y)}{\uparrow}}{2^n}$$

$$\Rightarrow S(\Gamma) = n 2^{n+1} \cdot 2^{-(n+1)} = n$$

$$l(\Gamma) \leq n$$

$$\Rightarrow \tau(e) \leq n (\log 2^n + \log \varepsilon^{-1})$$

Extensions to flow-based analysis

suppose we want to sample from the set of matchings in a graph, s.t.

$$\pi(M) = \frac{\mu^{|M|}}{Z} \quad \text{where } Z \text{ is the usual normalizing const.}$$

consider an MC that does -

- i) Choose an edge $e = (u, v)$ uniformly.
- ii) If $e \in M$ remove it with prob $\min(1, \mu^{-1})$
- iii) If $e \notin M$ and both u and v are unmatched add it with prob $\min(1, \mu)$.
- iv) If exactly one of u and v are matched remove it and add edge (u, v) .
- v) Else do nothing.

This is sort of a metropolis chain.

If X and Y differ by one edge -
and $|X| + 1 = |Y|$,

$$\pi(X) \cdot P(X, Y) = \frac{\mu^{|X|}}{Z} \cdot \frac{1}{m} \cdot \min(1, \mu)$$

$$\pi(Y) \cdot P(Y, X) = \frac{\mu^{|X|+1}}{Z} \cdot \frac{1}{m} \min(1, \mu^{-1})$$

if $\mu \leq 1$ (and similarly for $\mu > 1$).

$$\text{LHS} = \frac{\mu^{|X|}}{Z} \cdot \frac{1}{m} \cdot \mu = \frac{\mu^{|X|+1}}{Zm} = \text{RHS}.$$

If X and Y differ by 2 edges.
and $|X| = |Y|$

$$\pi(X) \cdot P(X, Y) = \frac{\mu^{|X|}}{Z} \cdot \frac{1}{m}$$

which are equal.

$$\pi(Y) \cdot P(Y, X) = \frac{\mu^{|Y|}}{Z} \cdot \frac{1}{m}$$

Unlike the hypercube example, we can't just count paths.

Let x, y be any two matchings in G . Consider $x \oplus y$ to be a graph with same vertex set but xor of edge sets. This must be a collection of alternating cycles and paths.

To define a path from x to y on the chain we take the first component of $x \oplus y$.

Here a "step" involves removing an edge from x and (if exists) adding the alternate edge to x . Repeat until all original edges from x

in the component is removed.

Now consider an MC transition (u, v) . How many paths from x to y pass through (u, v) .

Simultaneously consider a path from y to x . Now $y \oplus x$ is same as $x \oplus y$.

Now you can do some interesting stuff: explored in another paper to analyze this.

There is also a variant known as balanced flows.

Auxiliary Methods -

Comparison: modify MC without significantly changing the mixing time.

Let $\tilde{P}$ and P be reversible MCs on the same state space Ω with the same stat. dist π .

If $T_{\tilde{P}}(\epsilon)$ is known/bounded and we want $T_P(\epsilon)$ we use -

For each x, y with $\tilde{P}(x, y) > 0$, define γ_{xy} with $P(x_i, x_{i+1}) > 0$. Let $|\gamma_{xy}|$ be length of this path.

Let $\Gamma(z, w) = \{ (x, y) \mid (z, w) \in \gamma_{xy} \}$
the set of paths that use (z, w) transition.

Define $A = \max_{(z, w) \in E(P)} \left\{ \frac{1}{\pi(z) \cdot P(z, w)} \cdot \sum_{(x, y) \in \Gamma(z, w)} |\gamma_{xy}| \pi(x) \cdot \tilde{P}(x, y) \right\}$

Thm: $\text{Gap}(P) \geq \frac{1}{A} \text{Gap}(\tilde{P})$

Issue: same state space. very restrictive.

Decomposition

Idea:

Break down the chain into multiple pieces and relate overall runtime to those of the smaller chains.

Formally break $\Omega = \Omega_1, \dots, \Omega_m$ disjoint subsets.

Transitions are - $P_i(x, y) = p(x, y)$ if $x, y \in \Omega_i$
 $p_i(x, x) = 1 - \sum_{y \in \Omega_i} p(x, y)$.

$$\pi_i(A \subseteq \Omega) = \frac{\pi(A \cap \Omega_i)}{\pi(\Omega_i)}$$

Now transitions b/w partitions are determined by -

$$\bar{P}(i, j) = \frac{1}{\pi(\Omega_i)} \sum_{\substack{x \in \Omega_i \\ y \in \Omega_j}} \pi(x) p(x, y).$$

Thm: $\text{Gap}(P) \geq \frac{1}{2} \cdot \text{Gap}(\bar{P}) * \min_{i=1, \dots, m} \text{Gap}(P_i)$

Commonly Studied Models

1. Independent Sets: G be a graph and Ω be the

set of independent sets in G . Think of each indep set as a map from V to $\{0,1\}$ and no two neighbours can both be colored 1.

Let x_0, x_1 be the "desirability" of having a vertex in or out of an IS. Define weight-

$$w(I) = \prod_{v \in V} x_{f(v)}$$

This means $\Pi(I) = \frac{w(I)}{\sum_{I' \in \mathcal{I}} w(I')}$

2. Colorings

$$f: V \rightarrow \{1, \dots, k\}$$

Ω : set of proper k -colorings.

Let $x_0 \dots x_k$ be weights associated with each color.

$$w(C) = \prod_{v \in V} x_{f(v)}$$

3. Matchings

$$f: E \rightarrow \{0,1\}$$

Again x_0, x_1 are weights.

$$w(M) = \prod_{e \in E} x_{f(e)}$$

if $x_0 = 1, x_1 = u$ for integer u , $w(M)$ weights matchings as if G was a multigraph with u parallel

edges instead of each tree edge.

4. Pairwise influence models

For any n vertex $G(V, E)$, let $\Omega = \{1, \dots, q\}^n$ where $f: V \rightarrow \{1, \dots, q\}$

Define a symmetric set of weights $\{X_{ij} = X_{ji}\}$ for each pair $i, j \in \{1, \dots, q\}$ weight each configuration by

$$w(\sigma) = \prod_{(u,v) \in E} X_{f(u)f(v)}$$

Thus we can get q -colorings that favour some edges.

For ex if $X_{ij} = 1$ for all $i \neq j$ and $X_{ii} = 0$ for $i=j$ gives uniform distribution on proper k -coloring.

Unifying Framework for these problems

In statistical physics, models represent simple physical systems. Systems favor configurations that minimize energy.

The energy function on the space of configuration is determined by a Hamiltonian $H(\sigma)$.

$$w(\sigma) = \exp(-\beta H(\sigma)) \quad \text{where } \beta = \frac{1}{T}$$

For low values of β (high T) the diff b/w energies of configurations is damped.

At high β (low T) the diff are magnified.

Likelihood of each configuration: $\pi(\sigma) = \frac{w(\sigma)}{Z}$

$Z = \sum_{\sigma} w(\sigma)$ known as partition function.

This overall distr is called Gibbs/Boltzmann distr.

Let state space be the set of indep sets of a graph.

$$H(I) = - \sum_{v \in V} S_{v \in I} = -|I|, \text{ where } S = \begin{cases} 1 & \text{if } v \in I \\ 0, & \text{otherwise} \end{cases}$$

Now the prob distr -

$$\pi(I) = \frac{e^{-\beta H(I)}}{Z}$$

If $e^{\beta} = \lambda$, we get $\frac{\lambda^{|I|}}{Z}$. This is called the hard-core gas model.

Another model is the Ising model. Given a graph with n vertices, look at the 2^n ways of assigning spin ± 1 .

$$H(\sigma) = - \sum_{(u,v) \in E} \sigma_u \sigma_v = |D(\sigma)| - |A(\sigma)|$$

where D is the set of edges where $\sigma_u \neq \sigma_v$
 A $\sigma_u = \sigma_v$

$$\text{Then } \pi(\sigma) = \frac{1}{Z} \exp(-\beta(|D(\sigma)| - |A(\sigma)|))$$

$$= \frac{1}{Z} \exp(-\beta(|E| - 2|A(\sigma)|))$$

$$= \frac{\exp(\beta|E|) \cdot \exp(-2\beta|A(\sigma)|)}{\sum_{\sigma \in \{-1,1\}^n} \exp(\beta|E|) \cdot \exp(-2\beta|A(\sigma)|)}$$

$$= \frac{\exp(-2\beta |A(s)|)}{\sum_{\tau} \exp(-2\beta |A(\tau)|)}$$

This is a special case of pairwise influence where $q=2$, $x_{00} = x_{11} = \exp(-2\beta)$ and $x_{01} = x_{10} = 1$.

For larger q , the pairwise influence model specializes to the Potts model.

Same Terms "translated"

Physics

monomer-dimer coverings
 dimer coverings
 hard-core gas model
 spin
 ground states
 ground state in Potts
 Partition function
 connectivity
 activity / fugacity
 interaction
 ferromagnetism
 anti-ferromagnetism
 mean field
 Bethe lattice
 polynomial mixing
 rapid mixing

CS

Matchings
 perfect matchings
 Independent set
 bit
 Highest prob configurations
 vertex coloring
 Normalizing constant z .
 Degree
 vertex weights
 edge weights
 positive correlation
 negative correlation
 K_n
 complete regular free
 $O(n \log n)$ mixing.